

Rank Deficient Models

STA 721: Lecture 3

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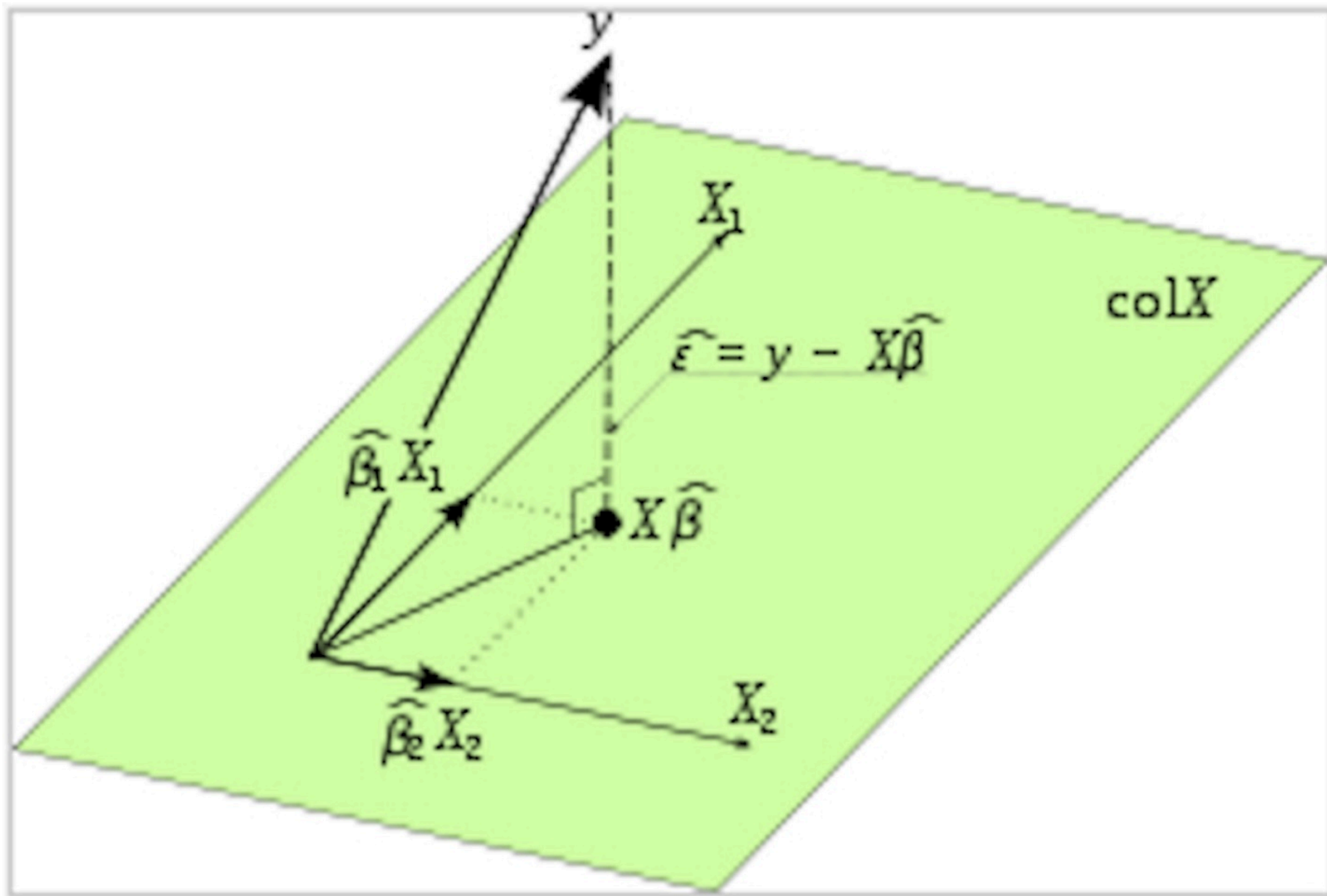
Outline

- Rank Deficient Models
- Generalized Inverses, Projections and MLEs/OLS
- Class of Unbiased Estimators

Readings: - Christensen Chapter 2 and Appendix B - Seber & Lee Chapter 3



Geometric View



Non-Full Rank Case

- Model: $\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\epsilon}$
- Assumption: $\boldsymbol{\mu} \in C(\mathbf{X})$ for $\mathbf{X} \in \mathbb{R}^{n \times p}$
- What if the rank of \mathbf{X} , $r(\mathbf{X}) \equiv r \neq p$?
- Still have result that the OLS/MLE solution satisfies

$$\mathbf{P}_{\mathbf{X}} \mathbf{Y} = \mathbf{X} \hat{\boldsymbol{\beta}}$$

- How can we characterize \mathbf{P} and $\hat{\boldsymbol{\beta}}$ in this case? 3 cases
 1. $p \leq n, r(\mathbf{X}) \neq p \Rightarrow r(\mathbf{X}) < p$
 2. $p > n, r(\mathbf{X}) \neq p$
 3. $p > n, r(\mathbf{X}) = p$

Focus on the first case for OLS/MLE for now...



Model Space

- $\mathcal{M} = C(\mathbf{X})$ is an r -dimensional subspace of \mathbb{R}^n
- \mathcal{M} has an $(n - r)$ -dimensional orthogonal complement \mathcal{N}
- each $\mathbf{y} \in \mathbb{R}^n$ has a unique representation as

$$\mathbf{y} = \hat{\mathbf{y}} + \mathbf{e}$$

for $\hat{\mathbf{y}} \in \mathcal{M}$ and $\mathbf{e} \in \mathcal{N}$

- $\hat{\mathbf{y}}$ is the orthogonal projection of \mathbf{y} onto \mathcal{M} and is the OLS/MLE estimate of $\boldsymbol{\mu}$ that satisfies

$$\mathbf{P}_{\mathbf{X}}\mathbf{y} = \mathbf{X}\hat{\boldsymbol{\beta}}$$

- $\mathbf{X}^T\mathbf{X}$ is not invertible so need another way to represent $\mathbf{P}_{\mathbf{X}}$ and $\hat{\boldsymbol{\beta}}$



Spectral Decomposition (SD)

Every symmetric $n \times n$ matrix, \mathbf{S} , has an eigen decomposition $\mathbf{S} = \mathbf{U}\mathbf{\Lambda}\mathbf{U}^T$

- $\mathbf{\Lambda}$ is a diagonal matrix with eigenvalues $(\lambda_1, \dots, \lambda_n)$ of \mathbf{S}
- \mathbf{U} is a $n \times n$ orthogonal matrix $\mathbf{U}^T\mathbf{U} = \mathbf{U}\mathbf{U}^T = \mathbf{I}_n$ ($\mathbf{U}^{-1} = \mathbf{U}^T$)
- the columns of \mathbf{U} form an Orthonormal Basis (ONB) for \mathbb{R}^n
- the columns of \mathbf{U} associated with non-zero eigenvalues form an ONB for $C(\mathbf{S})$
- the number of non-zero eigenvalues is the rank of \mathbf{S}
- the columns of \mathbf{U} associated with zero eigenvalues form an ONB for $C(\mathbf{S})^\perp$
- $\mathbf{S}^d = \mathbf{U}\mathbf{\Lambda}^d\mathbf{U}^T$ (matrix powers)



Positive Definite and Non-Negative Definite Matrices

▼ Definition: B.21 Positive Definite and Non-Negative Definite

A symmetric matrix \mathbf{S} is *positive definite* ($\mathbf{S} > 0$) if and only if $\mathbf{x}^T \mathbf{S} \mathbf{x} > 0$ for $\mathbf{x} \in \mathbb{R}^n$, $\mathbf{x} \neq \mathbf{0}_n$, and *positive semi-definite* or *non-negative definite* ($\mathbf{S} \geq 0$) if and only if $\mathbf{x}^T \mathbf{S} \mathbf{x} \geq 0$ for $\mathbf{x} \in \mathbb{R}^n$, $\mathbf{x} \neq \mathbf{0}_n$

Exercise

Show that a symmetric matrix \mathbf{S} is positive definite if and only if its eigenvalues are all strictly greater than zero, and positive semi-definite if all the eigenvalues are non-negative.



Projections

Let \mathbf{P} be an orthogonal projection matrix onto \mathcal{M} , then

1. the eigenvalues of \mathbf{P} , λ_i , are either zero or one
2. the trace of \mathbf{P} is the rank of \mathbf{P}
3. the dimension of the subspace that \mathbf{P} projects onto is the rank of \mathbf{P}
4. the columns of $\mathbf{U}_r = [u_1, u_2, \dots, u_r]$ form an ONB for the $C(\mathbf{P})$
5. the projection \mathbf{P} has the representation $\mathbf{P} = \mathbf{U}_r \mathbf{U}_r^T = \sum_{i=1}^r u_i u_i^T$ (the sum of r rank 1 projections)
6. the projection $\mathbf{I}_n - \mathbf{P} = \mathbf{I} - \mathbf{U}_r \mathbf{U}_r^T = \mathbf{U}_\perp \mathbf{U}_\perp^T$ where $\mathbf{U}_\perp = [u_{r+1}, \dots, u_n]$ is an orthogonal projection onto \mathcal{N}

MLE/OLS:

- $\mathbf{P}_X \mathbf{y} = \mathbf{U}_r \mathbf{U}_r^T \mathbf{y} = \mathbf{U}_r \tilde{\boldsymbol{\beta}}$
- Claim $\tilde{\boldsymbol{\beta}}$ is a MLE/OLS estimate of $\boldsymbol{\beta}$ where $\tilde{\mathbf{X}} = \mathbf{U}_r$.



Singular Value Decomposition & Connections to Spectral Decompositions

A matrix $\mathbf{X} \in \mathbb{R}^{n \times p}$, $p \leq n$ has a *singular value decomposition*

$$\mathbf{X} = \mathbf{U}_p \mathbf{D} \mathbf{V}^T$$

- \mathbf{U}_p is a $n \times p$ matrix with the first p eigenvectors in \mathbf{U} associated with the p largest eigenvalues of $\mathbf{X}\mathbf{X}^T = \mathbf{U}\mathbf{\Lambda}\mathbf{U}^T$ with $\mathbf{U}_p^T \mathbf{U}_p = \mathbf{I}_p$
- \mathbf{V} is a $p \times p$ orthogonal matrix associated with the p eigenvectors of $\mathbf{X}^T \mathbf{X} = \mathbf{V}\mathbf{\Lambda}_p \mathbf{V}^T$ where $\mathbf{\Lambda}_p$ is the diagonal matrix of eigenvalues associated with the p largest eigenvalues of $\mathbf{\Lambda}$
- $\mathbf{D} = \mathbf{\Lambda}_p^{1/2}$ are the singular values
- if \mathbf{X} has rank $r < p$, then $C(\mathbf{X}) = C(\mathbf{U}_p) = C(\mathbf{U}_r)$, where \mathbf{U}_r are the eigenvectors of \mathbf{U} or \mathbf{U}_p associated with the non-zero eigenvalues.



MLE/OLS for non-full rank case

- if $\mathbf{X}^T \mathbf{X}$ is invertible, $\mathbf{P}_X = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ and $\hat{\boldsymbol{\beta}}$ is the unique estimator that satisfies $\mathbf{P}_X \mathbf{y} = \mathbf{X} \hat{\boldsymbol{\beta}}$ or $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$
- if $\mathbf{X}^T \mathbf{X}$ is not invertible, replace \mathbf{X} by $\tilde{\mathbf{X}}$ that is rank r
- or represent $\mathbf{P}_X = \mathbf{X}(\mathbf{X}^T \mathbf{X})^- \mathbf{X}^T$ where $(\mathbf{X}^T \mathbf{X})^-$ is a generalized inverse of $\mathbf{X}^T \mathbf{X}$ and $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^- \mathbf{X}^T \mathbf{y}$



Generalized Inverses

▼ Definition: Generalized-Inverse (B.36)

A generalized inverse of any matrix \mathbf{A} : \mathbf{A}^- satisfies $\mathbf{A}\mathbf{A}^-\mathbf{A} = \mathbf{A}$

- A generalized inverse of \mathbf{A} symmetric always exists!

▼ Theorem: Christensen B.39

If \mathbf{G}_1 and \mathbf{G}_2 are generalized inverses of \mathbf{A} then $\mathbf{G}_1\mathbf{A}\mathbf{G}_2$ is also a generalized inverse of \mathbf{A}

- if \mathbf{A} is symmetric, then \mathbf{A}^- need not be!



Orthogonal Projections in General

Lemma B.43

If \mathbf{G} and \mathbf{H} are generalized inverses of $\mathbf{X}^T \mathbf{X}$ then

$$\begin{aligned}\mathbf{XG}\mathbf{X}^T\mathbf{X} &= \mathbf{XH}\mathbf{X}^T\mathbf{X} = \mathbf{X} \\ \mathbf{XG}\mathbf{X}^T &= \mathbf{XH}\mathbf{X}^T\end{aligned}$$

▼ Theorem: B.44

$\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-} \mathbf{X}^T$ is an orthogonal projection onto $C(\mathbf{X})$.

We need to show that (i) $\mathbf{P}\mathbf{m} = \mathbf{m}$ for $\mathbf{m} \in C(\mathbf{X})$ and (ii) $\mathbf{P}\mathbf{n} = \mathbf{0}$ for $\mathbf{n} \in C(\mathbf{X})^\perp$.

- i. For $\mathbf{m} \in C(\mathbf{X})$, write $\mathbf{m} = \mathbf{X}\mathbf{b}$. Then $\mathbf{P}\mathbf{m} = \mathbf{P}\mathbf{X}\mathbf{b} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-} \mathbf{X}^T \mathbf{X}\mathbf{b}$ and by Lemma B43, we have that $\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-} \mathbf{X}^T \mathbf{X}\mathbf{b} = \mathbf{X}\mathbf{b} = \mathbf{m}$
- ii. For $\mathbf{n} \perp C(\mathbf{X})$, $\mathbf{P}\mathbf{n} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-} \mathbf{X}^T \mathbf{n} = \mathbf{0}_n$ as $C(\mathbf{X})^\perp = N(\mathbf{X}^T)$.



MLEs & OLS

MLE/OLS satisfies

- $\mathbf{P}\mathbf{y} = \mathbf{X}\hat{\boldsymbol{\beta}}$
- $\mathbf{P}\mathbf{y} = \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-}\mathbf{X}^T\mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}\hat{\boldsymbol{\beta}}$ (does not depend on choice of generalized inverse)
- $\hat{\boldsymbol{\beta}} = (\mathbf{X}^T\mathbf{X})^{-}\mathbf{X}^T\mathbf{y}$
- $\hat{\boldsymbol{\beta}}$ is not unique - does depend on choice of generalized inverse unless \mathbf{X} is full rank



Moore-Penrose Generalized Inverse:

- Decompose symmetric $\mathbf{A} = \mathbf{U}\mathbf{\Lambda}\mathbf{U}^T$ (i.e $\mathbf{X}^T\mathbf{X}$)
- $\mathbf{A}_{MP}^- = \mathbf{U}\mathbf{\Lambda}^- \mathbf{U}^T$
- $\mathbf{\Lambda}^-$ is diagonal with

$$\lambda_i^- = \begin{cases} 1/\lambda_i & \text{if } \lambda_i \neq 0 \\ 0 & \text{if } \lambda_i = 0 \end{cases}$$

- Symmetric $\mathbf{A}_{MP}^- = (\mathbf{A}_{MP}^-)^T$
- Reflexive $\mathbf{A}_{MP}^- \mathbf{A} \mathbf{A}_{MP}^- = \mathbf{A}_{MP}^-$
- Can you construct another generalized inverse of $\mathbf{X}^T\mathbf{X}$?
- Can you find the Moore-Penrose generalized inverse of $\mathbf{X} \in \mathbb{R}^{n \times p}$?



Properties of OLS (full rank case)

How good is $\hat{\beta}$ as an estimator of β

- $\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X} \beta + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \epsilon$
- don't know ϵ , but can talk about behavior on average over
 - different runs of an experiment
 - different samples from a population
 - different values of ϵ
- with minimal assumption $\mathbf{E}[\epsilon] = \mathbf{0}_n$,

$$\begin{aligned} \mathbf{E}[\hat{\beta}] &= \mathbf{E}[(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{X} \beta + (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \epsilon] \\ &= (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{E}[\epsilon] \\ &= \beta \end{aligned}$$

- Bias of $\hat{\beta}$, $\text{Bias}[\hat{\beta}] = \mathbf{E}[\hat{\beta} - \beta] = \mathbf{0}_p$ - $\hat{\beta}$ is an unbiased estimator of β if $\mu \in C(\mathbf{X})$



Class of Unbiased Estimators

Class of linear statistical models:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

$$\boldsymbol{\epsilon} \sim P$$

$$P \in \mathcal{P}$$

An estimator $\tilde{\boldsymbol{\beta}}$ is unbiased for $\boldsymbol{\beta}$ if $\mathbf{E}_P[\tilde{\boldsymbol{\beta}}] = \boldsymbol{\beta} \quad \forall \boldsymbol{\beta} \in \mathbb{R}^p$ and $P \in \mathcal{P}$

Examples:

$$\mathcal{P}_1 = \{P = \mathbf{N}(\mathbf{0}_n, \mathbf{I}_n)\}$$

$$\mathcal{P}_2 = \{P = \mathbf{N}(\mathbf{0}_n, \sigma^2 \mathbf{I}_n), \sigma^2 > 0\}$$

$$\mathcal{P}_3 = \{P = \mathbf{N}(\mathbf{0}_n, \boldsymbol{\Sigma}), \boldsymbol{\Sigma} \in \mathcal{S}^+\}$$

\mathcal{P}_4 is the set of distributions with $\mathbf{E}_P[\boldsymbol{\epsilon}] = \mathbf{0}_n$ and $\mathbf{E}_P[\boldsymbol{\epsilon}\boldsymbol{\epsilon}^T] > \mathbf{0}$

\mathcal{P}_5 is the set of distributions with $\mathbf{E}_P[\boldsymbol{\epsilon}] = \mathbf{0}_n$ and $\mathbf{E}_P[\boldsymbol{\epsilon}\boldsymbol{\epsilon}^T] \geq \mathbf{0}$



Linear Unbiased Estimation

Exercise

1. Explain why an estimator that is unbiased for the model with parameter space $\beta \in \mathbb{R}^p$ and $P \in \mathcal{P}_{k+1}$ is unbiased for the model with parameter space $\beta \in \mathbb{R}^p$ and $P \in \mathcal{P}_k$.
2. Find an estimator that is unbiased for $\beta \in \mathbb{R}^p$ and $P \in \mathcal{P}_1$ that but is biased for $\beta \in \mathbb{R}^p$ and $P \in \mathcal{P}_2$.

Restrict attention to **linear** unbiased estimators

▼ Definition: Linear Unbiased Estimators (LUEs)

An estimator $\tilde{\beta}$ is a **Linear Unbiased Estimator (LUE)** of β if

1. linearity: $\tilde{\beta} = \mathbf{A}\mathbf{Y}$ for $\mathbf{A} \in \mathbb{R}^{p \times n}$
2. unbiasedness: $\mathbf{E}[\tilde{\beta}] = \beta$ for all $\beta \in \mathbb{R}^p$

- Are there other LUEs besides the OLS/MLE estimator?
- Which is “best”? (and in what sense?)

